Example ⁶⁹ : Opinion Polling Suppose that 40% of the population support a certain political candidate . Given a random sample of ¹⁵⁰ individuals, find 1.) the expected value and variance of the number of sampled individuals that favour the candidate . 2.) the probability that more than half the sample favour the candidate.

Example ⁶⁹ : Opinion Polling Suppose that 40% of the population support a certain polítical candidate. Given a random sample of ¹⁵⁰ individuals, find 1.) the expected value and variance of the number of sampled individuals that favour the candidate . 2.) the probability that more than half the sample favour the candidate. Let k_i be the answer of the i-th person, $\iota\iota$ yes $\overline{\nu}$ meaning 1 , and "no" meaning O. \Rightarrow $k_i \sim$ Bernoulli (p) with $\rho = 0.4$ Let $y := \sum_{i=1}^{n} x_i \implies y \sim Binom(u,p)$ $with u = 150$

$$
\Rightarrow E[y] = n \cdot \rho = 150 \times 0.4 = 60
$$

Var(y) = n \cdot \rho \cdot (1 - \rho) = 150 \times 0.4 \times 0.6 = 36
Check the rule of Hueub:
u \times \rho = 60 > S, u \times (1 - \rho) = 90 > S
= 7.646 m² + 10 m²

How can we compute this ?

1) Use the Binomial: Let ψ be the cdf of Binom (150, 0.4). R detivers $PLY > 753 = 1 - PLY = 753 = 1 - \psi(75)$

 $= 0.005225$

2) Approximate
$$
y
$$
 by a $y' \sim \sqrt{160.56}$
\n $P[y > 75] \approx P[y' > 75.5]^{\pi} = 1 - \Phi_{60,36} (75.5)$
\n= 0.004892 (with R)
\n* continuity correction
\n $1 - \Phi(\frac{75.5-60}{6})$

3) Approximation and Lookup in Z-Table Trewsform y' to ∂ \sim $\mathcal{N}(\partial_{\mathcal{U}})$: $P[y' > 75.5] = PL \frac{y' - 60}{c} > \frac{75.5 - 60}{6}$ \approx $P[E > \frac{75.5-60}{9}] = P[E > \frac{15.5}{6}]$ = $PIZ > 2.583$] = $1 - \Phi(2.585)$

 \approx 1 - 0.9951 = 0.0049

How Many Measurements are Needed ? We can use the CLT to determine the number of measurements needed for a required accuracy if we know the variance of the distribution of measurements .

 $\mathcal{L}(\mathcal{L})$ is a set of the set of $\mathcal{L}(\mathcal{L})$. In the set of the set of $\mathcal{L}(\mathcal{L})$

Example 66: *to* count to measure the distance to a

\nstar with

\n• **accuracy**
$$
a = A
$$
 (*i.e.*, with absolute error $\leq \frac{a}{2} = 0.5$) and

\n• **Corrary** $g = 4.5\%$.

\nThe variance of the measures is $3^2 - 2^2$.

\nLet d be the exact distance and 16 ; be the measures.

\nThe sample mean \overline{K}_n is close to a normal with

\nthe sample mean \overline{K}_n is close to a normal with

$$
\mu_n = \mu
$$
 and $\delta_n^2 = \frac{\delta^2}{n}$.

Then

 $\overline{\mathcal{X}}_{u}$ ple mean
 $\mu_n = \mu$
 $\overline{\mathcal{K}}_n - \mu_n$
 $\overline{\mathcal{S}}_n$ μ u χ u - E_{u} is close to a normal with
and $3u^{2} = \frac{3^{2}}{4}$.
 $= \frac{\widetilde{k}_{u} - \mu}{\delta/\sqrt{u}}$ a $W(o_{i}A)$ approximately. $\frac{1}{\sigma_{h}} = \frac{1}{\sigma l \sqrt{u}}$

We want a such that

$$
PL-\frac{a}{2} < \overline{E}_u - \mu < \frac{a}{2} \leq \gamma
$$

That is

$$
y \le P[-\frac{\pi}{\sigma} \frac{a}{2} < \frac{\pi}{\sigma}(\overline{x}_{u} - \mu) < \frac{\pi}{\sigma} \frac{a}{2}]
$$

\n $x P[-\frac{\pi}{\sigma} \frac{a}{2} < Z < \frac{\pi}{\sigma} \frac{a}{2}]$
\n $= 1 - 2(1-\Phi(\pi \frac{a}{2\sigma})) = 2 \cdot \Phi(\pi \frac{a}{2\sigma})) - 1$

hence

$$
\Phi(\sqrt{4} \frac{a}{23}) \ge \frac{1+8}{2}
$$
\n
$$
\Rightarrow \sqrt{4} \frac{a}{26} \ge \frac{-4}{2} (\frac{1+8}{2})
$$
\n
$$
\Leftrightarrow \sqrt{4} \ge \frac{26}{2} \overline{\Phi}^1(\frac{1+8}{2})
$$

This is an example where we need the inverse of the $cd\int$ to reason backward from a probasility to an argument.

We need an ⁿ such that

$$
\sqrt{4}
$$
 $\geq \frac{26}{a} \overline{\Phi}^{1}(\frac{1+\delta}{2})$

 w_1H_1

$$
a = 1, 3 = 2, \gamma = 0.95.
$$

This yields $\frac{1}{\sqrt{n}}$ \geq $\frac{2 \cdot 2}{1}$ $\frac{1}{2}$ ' $\left(\frac{4+0.95}{2}\right) = 4.\overline{\oplus}^1(0.975)$ $= 4 \times 1.960$ (in Z-table) $=$ 4 x 1.959964 (with R)

Hence $n \ge (4 \times 1.960)^{2} = 61.4656$

is ^a sufficiently large number of measurements.

4.2 Sample Variance

If we make measurements of some quantity, we consider this is evaluating ^a RV ^H. If we make several measurements, then we consider them as evaluations of a RUs \mathcal{H}_{n} . . $1/$ Hy that we washed, having the same distribution as k . How can we estimate the mean value of the distribution of H , i.e., $EIFJ$? The average $\overline{\mathcal{K}}_n$ of the \mathcal{K}_i , $\overline{\mathcal{K}}_n =$ - $\frac{1}{2}$ $\frac{1}{2}$ Ki , should be a good estimate .

How can we check that this is conceptually the right thing to do ?

Unbiased Estimators

Suppose $K_{1}, K_{1}, \ldots, K_{i}$, \cdots are i. E.d. RUS. A function $F(x_1,...,x_n)$, if applied to $\mathcal{X}_{n},...,\mathcal{X}_{n}$ defines a new random variable $F(k_{ij}...,k_{u})$. An example is k_{h} , which is defined by $F(x_1, ..., x_n) = \frac{\lambda}{\mu} (x_1 + ... + x_n) = \overline{x}_n$ Definition: Let \mathcal{H}_n , \mathcal{H}_n be i.i.d. RUS, $\overline{f}:\mathbb{R}^n\to\mathbb{R}$ a function and Obe a paremeles (Like mean, variance, or skew) of the distribution of the \mathcal{X}_i Then the bias of F with respect to θ for $(k_1,...,k_n)$ is $E(F(E_1,...,E_n)) - \theta$ and $F(E_{1},...,E_{n})$ is an unbiased estimator if the bies $\vec{\tau}$ s \vec{O} ,

EXAMPLES: (1) The average
$$
\overline{k}_{u} = \frac{1}{u} \sum_{i=1}^{u} k_{i}
$$
 is an unbiased
estimator of the mean μ .
(2) The average squared distance from the mean
 $\frac{1}{u} \sum_{i=1}^{u} (\overline{k}_{i} - \mu)^{2}$

is an unbiased estimator of the variance. (Note that we used je, not $\overline{\mathcal{E}}_{n}$.)

Proof: (1) If *have* calculated several times that

$$
E[\overline{X}_{u}] = \frac{1}{u} \sum_{i=1}^{u} E[\overline{X}_{i}] = \frac{1}{u} \cdot u \cdot \mu = \mu
$$

(2) Remember that $Var(X) = E[(X - \mu)^2]$. Thus

$$
E\left(\frac{1}{u}\sum_{i=1}^{u}(\mathcal{E}^{i}-\mu)^{2}\right)=\frac{1}{u}\sum_{i=1}^{u}\mathcal{E}\left(\mathcal{F}_{i}-\mu\right)^{2}=\frac{1}{u}\cdot u\cdot\sigma^{2}=\sigma^{2}
$$

Estimating the laurizon

\nConsider the function

\n
$$
\frac{1}{(k_1 - x_k)} = \frac{1}{n} \sum_{i=1}^{n} \frac{(k_i - \bar{x})^2}{(k_i - \bar{x})^2} \quad (*)
$$
\nIsif $k = \frac{1}{n} \sum_{i=1}^{n} x_i$.

\nThen our can calculate (see lecture works of 19110) that

\n
$$
E\left[\frac{1}{2} \left(\frac{k_i}{k_1} \ldots \frac{k_n}{k_n}\right)\right] = \frac{n - n}{n} \quad Var(k_1)
$$
\nThus, this is an estimator to find this.

\n
$$
\frac{n}{n-1} \int_0^2 (k_1 \ldots, k_n) dx = \frac{1}{n-1} \sum_{i=n}^{n} (k_i - \bar{x})^2 =: S^2
$$
\nis unbiased. This is also called the sample variable.

 x)" $7x^4$ is an abuse of notation, motivated by the attempt to estimate the variance