

We want to calculate:

$$\sum_{k=1}^{\infty} k \cdot \left(\frac{1}{2}\right)^k$$

It turns out that a more general sum is easier to calculate:

$$\sum_{k=1}^{\infty} k \cdot x^k, |x| < 1$$

Note that

$$k \cdot x^{k-1} = \frac{d}{dx} x^k$$

Consequently,

$$k x^k = x \cdot k \cdot x^{k-1} = x \cdot \frac{d}{dx} x^k$$

Reminder: Geometric series

$$\sum_{k=0}^{\infty} x^k = \frac{1}{1-x}, |x| < 1$$

Therefore, we can rewrite our sum as follows:

$$\begin{aligned}\sum_{k=1}^{\infty} k \cdot x^k &= \sum_{k=1}^{\infty} x \cdot k \cdot x^{k-1} = x \cdot \sum_{k=1}^{\infty} k \cdot x^{k-1} \\&= x \cdot \sum_{k=1}^{\infty} \frac{d}{dx} x^k \quad \left\{ \begin{array}{l} f' + g' = (f+g)' \\ \sum_{k=0}^{\infty} x^k = \frac{1}{1-x} \end{array} \right. \\&= x \cdot \frac{d}{dx} \sum_{k=1}^{\infty} x^k = x \cdot \frac{d}{dx} \frac{1}{1-x} \\&= x \cdot \frac{d}{dx} (1-x)^{-1} \quad \left\{ \begin{array}{l} \frac{d}{dx} x^\alpha = \alpha x^{\alpha-1} \\ \frac{d}{dx} (1-x)^{-1} = -1(1-x)^{-2} \end{array} \right. = \frac{x}{(1-x)^2}\end{aligned}$$

Now, we can plug in  $\frac{1}{2}$  for  $x$ :

$$\sum_{k=1}^{\infty} k \left(\frac{1}{2}\right)^k = \frac{\frac{1}{2}}{(1-\frac{1}{2})^2} = \frac{\frac{1}{2}}{\left(\frac{1}{2}\right)^2} = \frac{1}{\frac{1}{2}} = 2$$

So, the expected number of coin tosses until the first head is 2.

Definition Let  $X$  be a discrete R.V., with values  $x_1, \dots, x_n, \dots$

Then  $E[X] := \sum_{i=1}^n x_i P[X = x_i]$ , if  $X$  has  $n$  values

$E[X] = \sum_{i=1}^{\infty} x_i P[X = x_i]$ , if  $X$  has  $\infty$  many values,

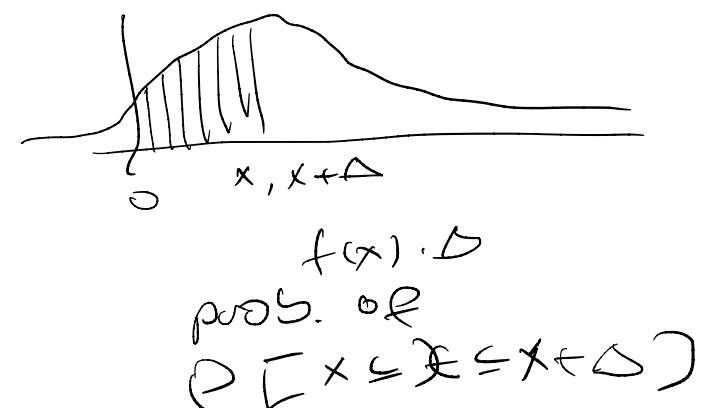
$E[X]$  is the expected value of  $X$ .

Definition Let  $X$  be a continuous R.V. with density  $f$ .

Then  $E[X] = \int_{-\infty}^{\infty} x f(x) dx$

is the expected value of  $X$

(if the integral exists)



Waiting time for arrival :

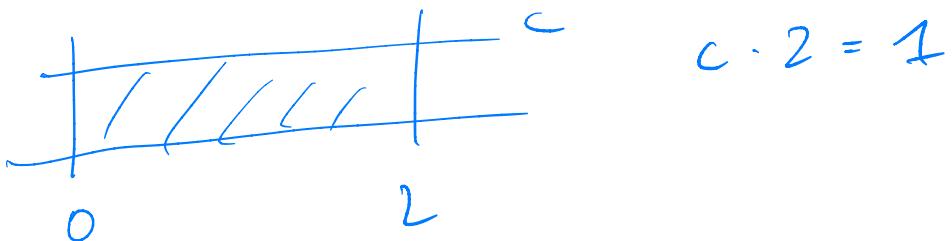
The waiting time  $R_U$ , takes value  $[0, 2]$

Density  $f(x) = \begin{cases} \frac{1}{2} & x \in [0, 2] \\ 0 & \text{otherwise} \end{cases}$

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx = \int_0^2 x \cdot \frac{1}{2} dx$$

$$= \left[ \frac{1}{2} \frac{x^2}{2} \right]_0^2 = \left[ \frac{x^2}{4} \right]_0^2 = \frac{2^2}{4} - \frac{0^2}{4} = \frac{4}{4} = 1$$

Uniform probability



## 2.2 Joint Distributions

Consider two RVs  $X, Y$  together. Study probabilities

$$P[X = x, Y = y]$$

or

$$P[a < X \leq b, c < Y \leq d]$$

Example 29: 9 batteries, 2 new, 3 part. charged, 4 empty.

Randomly select 3 out 9 batteries.

$$\begin{array}{ll} X & \# \text{ new batteries} \\ Y & \# \text{ partially charged} \end{array} \quad \begin{array}{l} X \in \{0, 1, 2\} \\ Y \in \{0, 1, 2, 3\} \end{array}$$

let  $p(x,y) = P[X=x, Y=y]$  joint pmf of  $X$  and  $Y$

$$p(0,0) = \frac{\binom{4}{3}}{\binom{9}{3}} = \frac{4}{84} \quad p(0,1) = \frac{\binom{3}{1} \binom{4}{2}}{\binom{9}{3}} = \frac{18}{84}$$

$$p(0,2) = \frac{\binom{3}{2} \binom{4}{1}}{\binom{9}{3}} = \frac{12}{84} \quad p(0,3) = \frac{\binom{3}{3}}{\binom{9}{3}} = \frac{1}{84}$$

$$p(1,0) = \frac{\binom{2}{1} \binom{4}{2}}{\binom{9}{3}} = \frac{12}{84}$$

$$\binom{9}{3} = \frac{9 \cdot 8 \cdot 7}{3 \cdot 2 \cdot 1} = 12 \cdot 7 = 84$$

$$p(1,1) = \frac{\binom{2}{1} \binom{3}{1} \binom{4}{1}}{\binom{9}{3}} = \frac{24}{84}$$

$$p(1,2) = \frac{\binom{2}{1} \binom{3}{2}}{\binom{9}{3}} = \frac{6}{84}$$

$$p(2,0) = \frac{\binom{2}{2} \binom{4}{1}}{\binom{9}{3}} = \frac{4}{84}$$

$$p(2,1) = \frac{\binom{2}{2} \binom{3}{1}}{\binom{9}{3}} = \frac{3}{84}$$

We have computed the joint pmf of  $X$  and  $Y$ .  
 Summarize in table (by multiples of  $\frac{1}{84}$ )

$x \backslash y$	0	1	2	3	Sum
0	4	18	12	1	35
1	12	24	6	0	42
2	4	3	0	0	7
Sum	20	45	18	1	84

Joint pmf

prob. of  $y = j$ , i.e.

$P[y = j]$

probability  
of  $X = i$ , i.e.  
 $P[X = i]$

marginal  
probabilities

Joint cumulative probability dist.

$$F(x, y) = P[X \leq x, Y \leq y]$$

Distribution of  $X$

$$\begin{aligned} F_X(x) &= P[X \leq x] = P[X \leq x, Y < \infty] \\ &= F(x, \infty) \end{aligned}$$

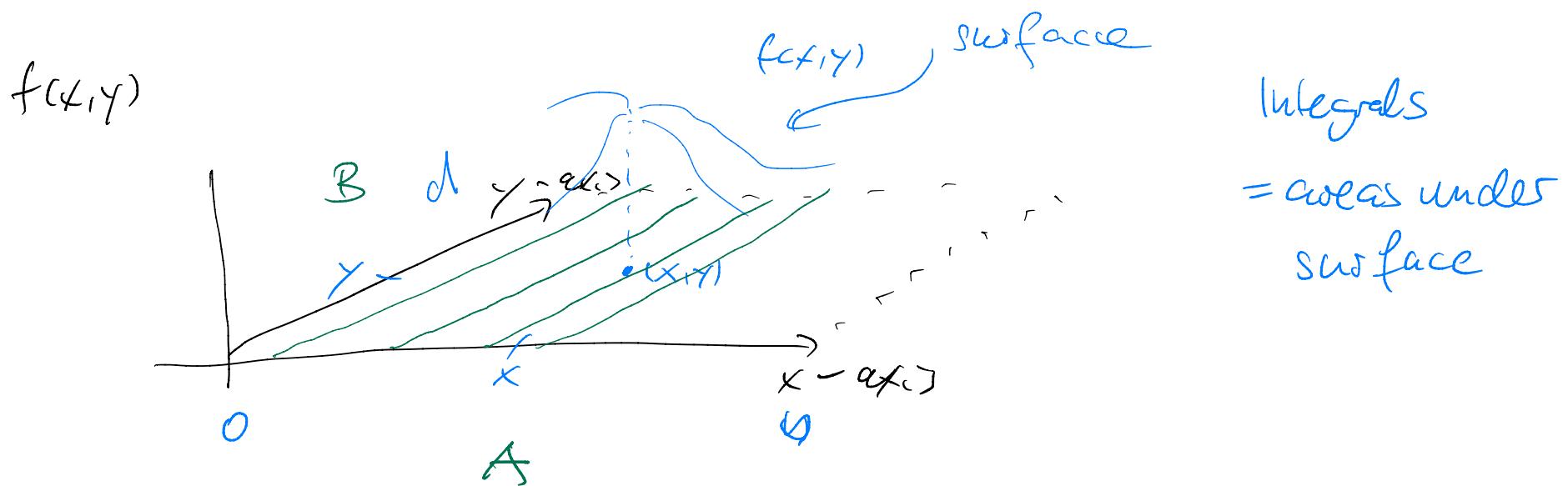
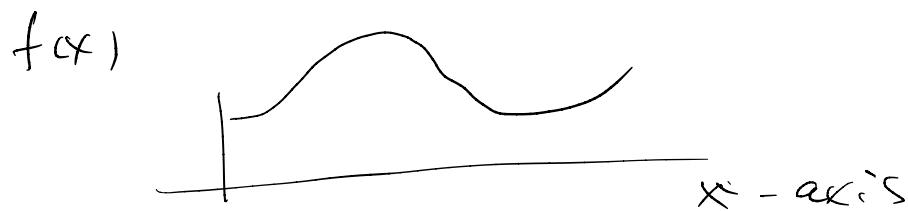
How we get the marginal pmf of  $X$  out of  
the joint pmf  $p(x, y) = P[X=x, Y=y]$ ?

$$p_x(x) = P[X=x] = \sum_{j=1}^n P[X=x, Y=y_j]$$
$$= \sum_{j=1}^n p(x, y_j)$$

$$p_y(y) = P[Y=y] = \sum_{i=1}^m p(x_i, y)$$

How can we model joint probabilities in cont. case?

discrete case:	joint prob	$P(x,y)$
continuous case:	joint pdf (p. density f.)	$f(x,y)$



Let  $X, Y$  continuous, joint pdf is a fct

$$f: \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R},$$

For  $C \subseteq \mathbb{R} \times \mathbb{R}$ , a reasonable subset, we have

$$P[(X, Y) \in C] = \iint_{(x,y) \in C} f(x, y) dx dy$$

Requirements of  $f$ :

$$f(x, y) \geq 0, \quad \iint_{\mathbb{R} \times \mathbb{R}} f(x, y) dx dy = 1$$

If  $A, B \subseteq \mathbb{R}$ , then

$$\begin{aligned} P[X \in A, Y \in B] &= \int_A \left( \int_B f(x, y) dy \right) dx \\ &= \int_B \left( \int_A f(x, y) dx \right) dy \end{aligned}$$

Example 32: Let the joint pdf of  $x, y$  be

$$f(x,y) = \begin{cases} 2e^{-x}e^{-2y} & x, y > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dx dy &= \int_0^{\infty} \int_0^{\infty} 2e^{-x}e^{-2y} dx dy \\ &= \int_0^{\infty} 2e^{-2y} \left( \int_0^{\infty} e^{-x} dx \right) dy = \int_0^{\infty} 2e^{-2y} \left[ -e^{-x} \right]_0^{\infty} dy \end{aligned}$$

$$\begin{aligned} \int_0^{\infty} 2e^{-2y} (0 - (-1)) dy &= \int_{-2.0}^{\infty} 2e^{-2y} dy = \left[ -e^{-2y} \right]_0^{\infty} \\ &= 0 - (-e^{-2 \cdot 0}) = 0 - (-1) = 1 \end{aligned}$$

$$P[X > 1, Y < 1] = ?$$

$$P[X < y] = ?$$

What are the marg. p.s of  $X$  and  $Y$ ?